

Probability & Statistics

L-18

Prob. Let X & Y have the joint pdf

$$p(x,y) = \exp\left(c + 4x + 4y - \frac{x^2}{2} - \frac{y^2}{2} - \frac{xy^2}{2}\right)$$

where c is a constant.

Q. Find $P_{X|Y}(x|y)$.

Ans.

$$P_y(y) = \int_{-\infty}^{\infty} p(x,y) dx$$

$$= \exp\left(c + 4y - \frac{y^2}{2}\right) \int_{-\infty}^{\infty} \exp\left(-\frac{(1+y^2)}{2}x^2 + 4x\right) dx$$

⋮

$$= \frac{\sqrt{2\pi}}{\sqrt{1+y^2}} \exp\left(c + 4y - \frac{y^2}{2} + \frac{8}{1+y^2}\right)$$

$$P_{X|Y}(x|y) = \frac{p(x,y)}{P_y(y)}$$

$$\sim N\left(\frac{4}{1+y^2}, \frac{1}{1+y^2}\right)$$

Observation: This is an example where conditional densities are normal but the joint pdf is not bivariate normal.

F-distribution & t-distribution.

Recall Gamma distribution.

$$p_X(x; \alpha, \beta) = \frac{\beta^\alpha (bx)^{\alpha-1} e^{-\beta x}}{\Gamma(\alpha)}$$

$$= \frac{\beta^\alpha x^{\alpha-1} e^{-\beta x}}{\Gamma(\alpha)}, \quad x \geq 0$$

my earlier defn.

$$p_X(x) = \frac{1}{\alpha^\alpha \Gamma(\beta)} x^{\beta-1} e^{-x/\alpha}$$

$$E(X^\alpha) = \frac{(\alpha+1)\dots\alpha}{\beta^\alpha}$$

$$\downarrow E(X) = \frac{\alpha}{\beta}, \quad E(X^2) = \frac{(\alpha+1)\alpha}{\beta^2}$$

$$\text{Var}(X) = \frac{\alpha}{\beta^2}.$$

Prop. Let X_1, X_2, \dots, X_n be a
set of independent random variables

$$X_1 \sim \Gamma(\alpha_1, \beta), \quad X_2 \sim \Gamma(\alpha_2, \beta), \dots, \quad X_n \sim \Gamma(\alpha_n, \beta)$$

$$\text{then } X_1 + X_2 + \dots + X_n \sim \Gamma(\alpha_1 + \alpha_2 + \dots + \alpha_n, \beta)$$

MGF of $\Gamma(\alpha, \beta)$

$$\begin{aligned}
 E(e^{tx}) &= \int_0^\infty e^{tx} \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x} dx \\
 &= \int_0^\infty \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-(\beta-t)x} dx \\
 &= \frac{\beta^\alpha}{(\beta-t)^\alpha} \int_0^\infty \frac{(\beta-t)^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-(\beta-t)x} dx \\
 &= \left(\frac{\beta}{\beta-t} \right)^\alpha.
 \end{aligned}$$

Calculate the mgf of $X_1 + X_2 + \dots + X_n$.

$$\begin{aligned}
 E(e^{t \sum_{j=1}^n x_j}) &= E\left(\prod_{j=1}^n e^{tx_j}\right) \\
 &= \prod_{j=1}^n \left(\frac{\beta}{\beta-t}\right)^{\alpha_j} \\
 &= \left(\frac{\beta}{\beta-t}\right)^{\sum \alpha_j}
 \end{aligned}$$

$$\therefore \sum_{j=1}^n X_j \sim \Gamma\left(\sum_{j=1}^n \alpha_j, \beta\right).$$

Recall: χ_n^2 distribution

$$\frac{1}{\Gamma\left(\frac{n}{2}, \frac{1}{2}\right)}$$

If X_i is i.i.d. standard normal then

$$X_1^2 + X_2^2 + \dots + X_n^2 \sim \chi_n^2.$$

Pf. Let $X \sim N(0, 1)$. Then

$$P(X^2 \leq x) \leftarrow \text{cdf of } X^2$$

$$= P(-\sqrt{x} < X < \sqrt{x})$$

$$= \int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt.$$

Then the pdf of X^2 is

$$p_{X^2}(x) = \frac{d}{dx} \int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$

Leibniz rule

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{(\sqrt{x})^2}{2}} (\sqrt{x})' - \frac{1}{\sqrt{2\pi}} e^{-\frac{(-\sqrt{x})^2}{2}} (-\sqrt{x})'$$

$$= \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{x}} e^{-x/2} = \frac{1}{\sqrt{2\pi}} x^{\frac{1}{2}-1} e^{-x/2} \sim \Gamma\left(\frac{1}{2}, \frac{1}{2}\right)$$

F-distribution

Let us consider two independent random variables

$$X \sim \chi^2_k = \Gamma\left(\frac{k}{2}, \frac{1}{2}\right)$$

$$Y \sim \chi^2_m = \Gamma\left(\frac{m}{2}, \frac{1}{2}\right)$$

Then the distribution of the random variable

$$Z = \frac{X/k}{Y/m} \quad \text{is called}$$

Fisher - distribution with degrees of freedom k & m , denoted by $F_{k,m}$.

The pdf is given by

$$\frac{\Gamma\left(\frac{k+m}{2}\right)}{\Gamma\left(\frac{k}{2}\right) \Gamma\left(\frac{m}{2}\right)} t^{\frac{k}{2}-1} (1+t)^{-\frac{k+m}{2}}$$

Student's t-distribution

Consider the random variable

$$X_1$$

$$T = \frac{Y_1}{\sqrt{\frac{1}{n} (Y_1^2 + Y_2^2 + \dots + Y_n^2)}}$$

where $X_1, Y_1, Y_2, \dots, Y_n$ are i.i.d. ~~not~~ standard normal.

The pdf of T is given by

$$f_T(t) = \frac{\Gamma(\frac{n+1}{2})}{\Gamma(\frac{1}{2}) \Gamma(\frac{n}{2})} \frac{1}{\sqrt{n}} \left(1 + \frac{t^2}{n}\right)^{-\frac{n+1}{2}}$$