

# Probability and Statistics

## MA20205

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# Descriptive statistics

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- 2 These functions help us predicting the behavior of the random variable corresponding to a real life problem
- 3 There are many parameters, mainly the mean ( $\mu$ ) and variance  $\sigma^2$  characterize the random variable
- 4 In practice, however the pdf or cdf are not known and hence the parameters are also not known
- 5 Goal: how to determine a reasonable pdf and approximate the values for the distribution parameters from a data set

# Sampling distributions

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**Population** The (large) set of objects about which some inferences are to be made is called the population.

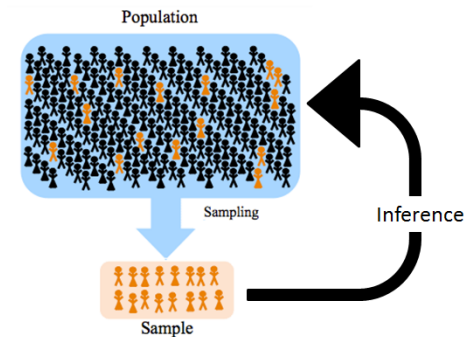
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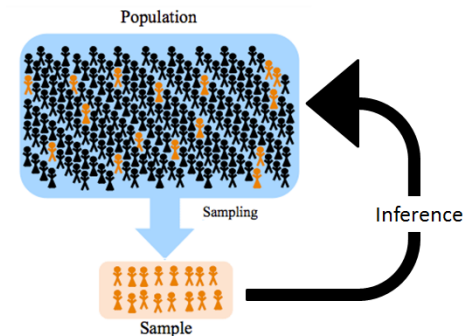
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**Population** The (large) set of objects about which some inferences are to be made is called the population. there should be at least one random variable relative to the population whose behavior is to be studied. Since we can not inspect all the elements of the population, we have to select a sample of objects from the population.

# Sampling distributions



# Sampling distributions



**Random sampling** Select  $n$  objects from the population such that the selection of one object neither ensures nor precludes the selection of any other. Thus the selection of one object is independent of selection of any other. This set is called **random sample**.

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## Random sample

A random of size  $n$  from the distribution of  $X$  is a collection of  $n$  independent random variables, each with the same distribution of  $X$

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A statistic is a random variable whose numerical value can be determined from a random sample  $X_1, \dots, X_n$ . Some of the important statistics for statisticians are  $\sum_{i=1}^n X_i$ ,  $\sum_{i=1}^n X_i^2$ ,  $\max_i X_i$ ,  $\min_i X_i$ , sample mean, sample variance etc.

## Statistics

① Sample mean:  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

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The statistic  $S = \sqrt{S^2}$  is called the **sample standard deviation**

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**Question** Why  $1/(n-1)$  instead of  $1/n$  in the definition of  $S^2$ ?

Then  $S^2$  becomes an unbiased (will be defined later) statistic for  $\sigma_X^2$ .

# Sampling distributions

We denote  $\bar{x}$  and  $s^2$  as realizations from a particular sample. We denote  $X$  as the population random variable. Then note that  $\mu_X$  and  $\bar{X}$  need not be same.

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Problems..

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If  $X_1, \dots, X_n$  are mutually independent random variables with respective means  $\mu_1, \dots, \mu_n$  and variances  $\sigma_1^2, \dots, \sigma_n^2$  then the mean and variance of  $Y = \sum_{i=1}^n a_i X_i$ ,  $a_i \in \mathbb{R}$  are given by

$$\mu_Y = \sum_{i=1}^n a_i \mu_i \text{ and } \sigma_Y^2 = \sum_{i=1}^n a_i^2 \sigma_i^2.$$



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Proof

$$\mu_Y = E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E(X_i) = \sum_{i=1}^n a_i \mu_i$$

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Since  $X_i$ s are mutually independent,  $\text{Cov}(X_i, X_j) = 0$  when  $i \neq j$ . Thus

$$\sigma_Y^2 = \text{Var}(Y) = \sum_{i=1}^n \text{Var}(a_i X_i) = \sum_{i=1}^n a_i^2 \text{Var}(X_i) = \sum_{i=1}^n a_i^2 \sigma_i^2.$$

## Sampling distributions

If  $X_1, \dots, X_n$  are independent random variables with respective mgfs  $M_{X_i}(t)$ ,  $i = 1, \dots, n$  then the mgf of  $Y = \sum_{i=1}^n a_i X_i$  is given by

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If  $X_1, \dots, X_n$  are independent random variables with  $X_i \sim \mathcal{N}(\mu_i, \sigma_i^2)$  then the random variable  $Y = \sum_{i=1}^n a_i X_i \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$ , where

$$\mu_Y = \sum_{i=1}^n a_i \mu_i \text{ and } \sigma_Y^2 = \sum_{i=1}^n a_i^2 \sigma_i^2$$

(the proof from above result)

## Sampling distributions

If  $X_1, \dots, X_n$  is a random sample of size  $n$  from a normal distribution with mean and variance  $\sigma^2$  then  $\bar{X} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$  (proof follows from above result)

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If  $X_1, \dots, X_n$  are independent random variables with respective distributions  $\chi^2(r_1), \dots, \chi^2(r_n)$  then

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**Proof** Since  $X_i \sim \chi^2(r_i)$ , the mgf of  $X_i$  is

$$M_{X_i}(t) = (1 - 2t)^{-\frac{r_i}{2}}.$$

Then from the above result regarding mgf,

$$M_Y(t) = \prod_{i=1}^n M_{X_i}(t) = \prod_{i=1}^n (1 - 2t)^{-\frac{r_i}{2}} = (1 - 2t)^{-\frac{1}{2} \sum_{i=1}^n r_i}$$



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Before we discuss our next result, let us denote

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \text{ and } S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

## Sampling distributions

If  $X_1, \dots, X_n$  is a random sample of size  $n$  from the normal distribution  $\mathcal{N}(\mu, \sigma^2)$  then

- 1  $\frac{(n-1)S_n^2}{\sigma^2} \sim \chi^2(n-1)$  and
- 2  $\bar{X}_n$  and  $S_n^2$  are independent

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**Proof** The proof follows by induction. let us prove it for  $n = 2$ . Since  $X_i \sim \mathcal{N}(\mu, \sigma^2)$ ,  $i = 1, \dots, n$  then  $X_1 + X_2 \sim \mathcal{N}(2\mu, 2\sigma^2)$  and  $X_1 - X_2 \sim \mathcal{N}(0, 2\sigma^2)$ . Hence

$$\frac{X_1 - X_2}{2\sigma^2} \sim \mathcal{N}(0, 1)$$

and thus

$$\frac{(X_1 - X_2)^2}{2\sigma^2} \sim \chi^2(1)$$

which proves  $S_2^2 \sim \chi^2(1)$ .

# Sampling distributions

Now, since  $X_1$  and  $X_2$  are independent,

$$\begin{aligned} & \text{Cov}(X_1 + X_2, X_1 - X_2) && (1) \\ = & \text{Cov}(X_1, X_1) + \text{Cov}(X_1, X_2) - \text{Cov}(X_2, X_1) - \text{Cov}(X_2, X_2) \\ = & \sigma^2 + 0 - 0 - \sigma^2 = 0 \end{aligned}$$

Thus  $X_1 + X_2$  and  $X_1 - X_2$  are uncorrelated bivariate normal random variables. This yields that they are independent. Therefore  $\frac{1}{2}(X_1 + X_2)$  and  $\frac{1}{2}(X_1 - X_2)^2$  are independent i.e.  $X_2$  and  $S_2^2$  are independent.

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Suppose  $X$  is a nonnegative random variable with variance  $\sigma^2$ . Then

$$P(|X - \mu| \geq t) \leq \frac{\sigma^2}{t^2}$$

for all  $t > 0$ .

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### Definition

Suppose  $X_1, X_2, \dots$  is a sequence of random variables on a sample space. Then the sequence **converges in probability** to the random variable  $X$  if, for any  $\epsilon > 0$ ,

$$\lim_{n \rightarrow \infty} P(|X_n - X| < \epsilon) = 1.$$



# Sampling distributions

## Weak law of large numbers (WLLN)

Let  $X_1, X_2, \dots$  be a sequence of iid random variables with  $\mu = E(X_i)$  and  $\sigma^2 = \text{Var}(X_i) < \infty$  for  $i = 1, 2, \dots$ . Then

$$\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| \geq \epsilon) = 0 \text{ i.e. } \lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| < \epsilon) = 1.$$

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**Proof** We proved before that  $E(\bar{X}_n) = \mu$  and  $\text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$ . Then by Chebyshev's inequality,

$$P(|\bar{X}_n - \mu| \geq \epsilon) \leq \frac{\sigma^2}{n\epsilon^2}$$

for  $\epsilon > 0$ . Then the result follows by setting  $n \rightarrow \infty$  both sides.

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For instance, consider an experiment of tossing a coin infinitely many times. Let  $X_i$  be 1 if the  $i$ th toss is head, and 0 otherwise. Then WLLN says

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n} \rightarrow \frac{1}{2} \text{ as } n \rightarrow \infty$$

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However, there can exist a sequence of coin tosses like

*HHHHHHHHHHH...*

but the WLLN says that the probability of occurrence of such a sequence is zero.

# Sampling distributions

## Definition

Let  $X_1, X_2, \dots$  be a sequence of random variables on a sample space  $S$ . Then the sequence  $X_n(\omega)$  **converges almost surely** to  $X(\omega)$  if

$$P\left(\left\{\omega \in S \mid \lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)\right\}\right) = 1,$$

where  $X$  is a random variable on the sample space  $S$ .

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## Strong law of large numbers (SLLN)

Let  $X_1, X_2, \dots$  be a sequence of iids with  $\mu = E(X_i), i = 1, 2, \dots$ . Then

$$P\left(\lim_{n \rightarrow \infty} \bar{X}_n = \mu\right) = 1.$$



# Sampling distributions

**Recall:** Let  $X_1, \dots, X_n$  be a random sample. Then these are iids with a common pdf which is the pdf of the population. Further, if the population pdf is normal then the the sample mean is normal i.e. if  $X_i$  is from the distribution

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

then

$$\bar{X}_n \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

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$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

then

$$\bar{X}_n \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

# Sampling distributions

## Central limit theorem/Lindeberg-Levy theorem

Let  $X_1, X_2, \dots, X_n$  be a random sample from a distribution with mean  $\mu$  and variance  $\sigma^2 < \infty$ . Then the limiting distribution of

$$Z_n = \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}}$$

is standard normal, i.e.  $Z_n$  converges in distribution to the standard normal random variable.