

Probability and Statistics

MA20205

Bibhas Adhikari

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Lecture 15
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Bivariate normal distribution

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$$f(x, y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} e^{-\frac{1}{2}Q(x,y)}$$

where $\mu_1, \mu_2 \in \mathbb{R}$, $\sigma_1, \sigma_2 \in (0, \infty)$ and $\rho \in (-1, 1)$ are parameters, and

$$Q(x, y) = \frac{1}{1-\rho^2} \left[\left(\frac{x-\mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x-\mu_1}{\sigma_1} \right) \left(\frac{y-\mu_2}{\sigma_2} \right) + \left(\frac{y-\mu_2}{\sigma_2} \right)^2 \right]$$

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Denote:

$$(X, Y) \sim \mathcal{N}(\mu_1, \mu_2, \sigma_1, \sigma_2, \rho)$$

Bivariate normal/Gaussian distribution

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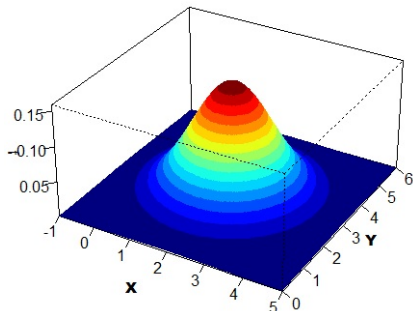
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- 3 σ_1^2 and σ_2^2 measure the spread of the mountain in the x -direction and y -direction respectively
- 4 ρ determines the shape and orientation

Watch: [The link](#)



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$$f_Y(y) = \frac{1}{\sigma_2\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{y-\mu_2}{\sigma_2}\right)^2}$$

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Remark However the converse need not be true!!

Bivariate normal/Gaussian distribution

proof of marginal distribution

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$$Q(x, y) = \frac{1}{1 - \rho^2} \left[\left(\frac{x - \mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x - \mu_1}{\sigma_1} \right) \left(\frac{y - \mu_2}{\sigma_2} \right) + \left(\frac{y - \mu_2}{\sigma_2} \right)^2 \right]$$

Bivariate normal/Gaussian distribution

proof of marginal distribution

$$\begin{aligned} Q(x, y) &= \frac{1}{1 - \rho^2} \left[\left(\frac{x - \mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x - \mu_1}{\sigma_1} \right) \left(\frac{y - \mu_2}{\sigma_2} \right) \right. \\ &\quad \left. + \left(\frac{y - \mu_2}{\sigma_2} \right)^2 \right] \\ &= \frac{1}{1 - \rho^2} \left[\left(\frac{x - \mu_1}{\sigma_1} - \rho \frac{y - \mu_2}{\sigma_2} \right)^2 + (1 - \rho^2) \left(\frac{y - \mu_2}{\sigma_2} \right)^2 \right] \\ &= \frac{(x - a)^2}{(1 - \rho^2)\sigma_1^2} + \frac{(y - \mu_2)^2}{\sigma_2^2} \end{aligned}$$

where $a = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (y - \mu_2)$.

Bivariate normal/Gaussian distribution

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where $a = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (y - \mu_2)$. Hence

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \frac{1}{\pi \sigma_1 \sigma_2 \sqrt{1 - \rho^2}} e^{-\frac{(y - \mu_2)^2}{2\sigma_2^2}} \int_{-\infty}^{\infty} e^{-\frac{(x - a)^2}{2(1 - \rho^2)\sigma_1^2}} dx$$

Bivariate normal/Gaussian distribution

If $(X, Y) \sim \mathcal{N}(\mu_X, \mu_Y, \sigma_X, \sigma_Y, \rho)$ then

$$E(X) = \mu_X, E(Y) = \mu_Y, \text{Var}(X) = \sigma_X^2, \text{Var}(Y) = \sigma_Y^2$$

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$$\text{Correlation coefficient} = \rho, M(s, t) = e^{\mu_X s + \mu_Y t + \frac{1}{2}(\sigma_X^2 s^2 + 2\rho\sigma_X\sigma_Y st + \sigma_Y^2 t^2)}$$

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Proof $X \sim \mathcal{N}(\mu_X, \sigma_X^2), Y \sim \mathcal{N}(\mu_Y, \sigma_Y^2)$. Then

$W = sX + tY \sim \mathcal{N}(\mu_W, \sigma_W^2)$ where

$$\mu_W = s\mu_X + t\mu_Y, \sigma_W^2 = s^2\sigma_X^2 + 2st\rho\sigma_X\sigma_Y + t^2\sigma_Y^2.$$

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Therefore, the mgf of W is $M(\tau) = e^{\mu_W \tau + \frac{1}{2}\tau^2 \sigma_W^2}$.

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Therefore, the mgf of W is $M(\tau) = e^{\mu_W\tau + \frac{1}{2}\tau^2\sigma_W^2}$. Then mgf of (X, Y) is

$$\begin{aligned} M(s, t) &= E(e^{sX+tY}) = e^{\mu_W + \frac{1}{2}\sigma_W^2} \\ &= e^{\mu_X s + \mu_Y t + \frac{1}{2}(\sigma_X^2 s^2 + 2\rho\sigma_X\sigma_Y st + \sigma_Y^2 t^2)} \end{aligned}$$

Bivariate normal/Gaussian distribution

Let $f(x, y)$ be the joint pdf of (X, Y) . Then conditional density of Y given $X = x$ is

$$f_{Y|X}(y|x) = \frac{1}{\sigma_Y \sqrt{2\pi(1-\rho^2)}} e^{-\frac{1}{2} \left(\frac{y-b}{\sigma_Y \sqrt{1-\rho^2}} \right)^2}$$

where

$$b = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X).$$

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where

$$b = \mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x - \mu_X).$$

Similarly,

$$f_{X|Y}(x|y) = \frac{1}{\sigma_X \sqrt{2\pi(1-\rho^2)}} e^{-\frac{1}{2} \left(\frac{x-c}{\sigma_X \sqrt{1-\rho^2}} \right)^2}$$

where

$$c = \mu_X + \rho \frac{\sigma_X}{\sigma_Y} (y - \mu_Y)$$

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If $(X, Y) \sim \mathcal{N}(\mu_X, \mu_Y, \sigma_X, \sigma_Y, \rho)$ then

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$$\text{Var}(Y|x) = \sigma_Y^2 (1 - \rho^2)$$

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$$E(X|y) = \mu_X + \rho \frac{\sigma_X}{\sigma_Y} (y - \mu_Y)$$

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$$E(X|y) = \mu_X + \rho \frac{\sigma_X}{\sigma_Y} (y - \mu_Y)$$

$$\text{Var}(Y|x) = \sigma_Y^2 (1 - \rho^2)$$

$$\text{Var}(X|y) = \sigma_X^2 (1 - \rho^2)$$

Interesting result (Cramer, 1941)

Two random variables X and Y have a joint bivariate normal distribution if and only if every linear combination of X and Y has univariate normal distribution