# Big Data Analysis (MA60306)

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Spring 2022-23, IIT Kharagpur

Lecture 21 March 22, 2023

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Generation of uniformly distributed random variables and methods from transforming it to other distributions

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Relevance to cryptography - dynamic concept of randomness

 $\rightarrow$  a process is "random" if the known conditional probability of the next event, given the previous history (or any other information, for that matter) is no different from the known unconditional probability.

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- $\rightarrow\,$  This kind of definition leads to the concept of a "one-way function"
- $\rightarrow$  A one-way function is a function f, such that for any x in its domain, f(x) can be computed in polynomial time, and given f(x), x cannot be computed in polynomial time

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- $\rightarrow\,$  In random number generation, the function of interest yields a stream of "unpredictable" numbers, that is,

$$x_i = f(x_{i-1}, x_{i-2}, \ldots, x_{i-k})$$

is easily computable; but  $x_{i-1}$ , given  $x_i, x_{i-2}, ..., x_{i-k}$ , is not easily computable

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(pseudo)Random number generation - a mechanism for generating a sequence of random variables  $U_1, U_2, \ldots$  with the property that

each  $U_i$  is uniformly distributed between 0 and 1

the  $U_i$  are mutually independent i.e. the value of  $U_i$  should not be predictable from  $U_1, \ldots, U_{i-1}$ 

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Goal produce a finite sequence of numbers  $u_1, \ldots, u_K$  in the unit interval Uniformity - If the K is large then the fraction of values falling in any subinterval of the unit interval should be approximately the length of the subinterval

Independence - there should be no discernible pattern among the values i.e. statistical test for independence should not easily reject segments of the sequence  $u_1, \ldots, u_K$ 

Modular arithmetic Two integers a, b are said to be congruent or equivalent modulo m if a - b is divisible by m, we write it as

 $a \equiv b \mod m$ 

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Reduction of *b* modulo *m* can be defined as:

$$a=b-\lfloor b/m\rfloor m.$$

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Linear (multiplicative) congruential generator

$$\begin{array}{rcl} x_{i+1} &=& ax_i \bmod m \\ u_{i+1} &=& x_{i+1}/m \end{array} \tag{1}$$

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the integers a and m determine the values generated, given an initial value  $1 \le x_0 \le m - 1$ , specified by the user.

If a, m are properly chosen then  $u_i$ 's 'look like' they are randomly and uniformly distributed. The recurrence relation (1) is equivalent to the recurrence

$$u_i \equiv a u_{i-1} \mod 1$$
 with  $0 < u_i < 1$ 

Note that the operation

$$y \mod m = y - \left\lfloor \frac{y}{m} \right\rfloor m$$

For example,  $7 \mod 5 = 2$ .

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$$1, 6, 3, 7, 9, 10, 5, 8, 4, 2, 1, 6, \ldots$$

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 $\rightarrow$  (Homework) What is your observation for different choices of  $x_0$ ?

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General considerations

→ Period length - The generator of the above form eventually repeat itself. The longest possible period for a linear congruent generator with *mod m* is m - 1, and with full period the gaps between the values  $u_i$  are 1/m. Thus the larger *m* is more closely the values can approximate a uniform distribution

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Full period - A linear congruential generator is said to have full period if it produces all m - 1 distinct values before repeating

(D. H. Lehmer, 1948) A general linear (mixed) congruential generator:

$$x_{i+1} = (ax_i + c) \mod m$$
  
 $u_{i+1} = x_{i+1}/m$  (2)

a, m, c are integers (a is called 'multiplier, c is called the 'increment', and m is called the 'modulus' of the generator)

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If *m* is a power of 2, the generator has full period if *c* is odd and a = 4n + 1 for some integer *n* 

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Question How large should the value of m be?

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Modulus <i>m</i>	Multiplier a
$2^{31} - 1$	16807
(= 2147483647)	
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2147483399	1226874159 40692

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Question How large should the value of m be? Literature:

Modulus <i>m</i>	Multiplier a
$2^{31} - 1$	16807
(= 2147483647)	
	1226874159
2147483399	1226874159 40692

Note Currently the numbers used as moduli in production random number generators are often primes in particular, Mersenne primes, which have the form  $2^p - 1$ . Numbers of this form for  $p \le 31$  are prime except for the three values: p = 11, 23, and 29. Most larger values of p do not yield prime (Write a program and check!), however p = 859433 does give a prime.

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Tests of linear congruential generator form any transformation on subsequences of a produced sequence of numbers, determine the distribution of the transformation under the null hypothesis of independent uniformity of the sequence, and perform a goodness-of-fit test of that distribution

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#### Some other methods

Multiple Recursive Generators Lagged Fibonacci Inversive Congruential Generators Nonlinear Congruential Generators Matrix Congruential Generators and many more including Monte Carlo methods

Bibhas Adhikari (Spring 2022-23, IIT Kharag

Sampling from a nonuniform distribution - usually done by applying a transformation to uniform sampler or from a sequence of uniform samplers, other methods use a random walk sequence, a Markov chain

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The performance of the algorithms is judged by - in speed, in accuracy, in storage requirements, and in complexity of coding

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Now we assume that there is a good way to generate independent samples of the uniform random variable U on the interval (0, 1)

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Sampling finite and discrete rvs

Bernoulli random variable: If

$$X = egin{cases} 1 ext{ if } U \leq p \ 0 ext{ otherwise} \end{cases}$$

then  $X \sim Ber(p)$  since 1 will be sampled with probability p, and 0 will be sampled with probability 1 - p.

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Inverse transform technique: Let  $X = \{x_1, \ldots, x_n\}$  be a random variable with probability distribution p, and where  $x_1 \leq \ldots \leq x_n$ . Then define

$$q_i = P(X \le x_i) = \sum_{j=1}^{l} p(x_j)$$

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Inverse transform technique: Let  $X = \{x_1, \ldots, x_n\}$  be a random variable with probability distribution p, and where  $x_1 \leq \ldots \leq x_n$ . Then define

$$q_i = P(X \le x_i) = \sum_{j=1}^i p(x_j)$$

Then the following is the sampling formula for X as follows:

$$X = \begin{cases} x_1 \text{ if } U < q_1 \\ x_2 \text{ if } q_1 \le U < q_2 \\ \vdots \vdots \\ x_{n-1} \text{ if } q_{n-2} \le U < q_{n-1} \\ x_n \text{ otherwise} \end{cases}$$

Note that  $X = x_i$  in the event that  $q_{i-1} \le U < q_i$ , which has probability  $p = q_i - q_{i-1} = p(x_i)$ .

Geometric rv For the uniform distribution U,

$$X = \left\lfloor \frac{\ln U}{\ln q} \right\rfloor + 1 \sim G(p)$$

with parameter p = 1 - q

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Geometric rv For the uniform distribution U,

$$X = \left\lfloor \frac{\ln U}{\ln q} \right\rfloor + 1 \sim G(p)$$

with parameter p = 1 - qProof First sample U(0, 1) and then return k, where

$$\sum_{n=1}^{k-1} (1-p)^{n-1} p \leq U < \sum_{n=1}^k (1-p)^{n-1} p$$

which implies

$$1 - (1 - p)^{k - 1} \le U < 1 - (1 - p)^k \Rightarrow (1 - p)^k < 1 - U \le (1 - p)^{k - 1}$$

using

$$\sum_{n=1}^{k} ar^{n-1} = a \frac{r^k - 1}{r-1}$$

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Taking log both side and dividing by the negative number ln(1-p) then

$$k-1 \leq \frac{\ln(1-U)}{\ln(1-p)} < k \Rightarrow k = \left\lfloor \frac{\ln(1-U)}{\ln(1-p)} \right\rfloor + 1.$$

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Finally, setting q = 1 - p, and noting that 1 - U is also uniformly distributed over [0, 1, ] we have

$$k = \left\lfloor \frac{\ln U}{\ln q} \right\rfloor + 1$$

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Continuous rv - Inverse Transform Method Let X be a continuous random variable with cdf F(x) which possesses an inverse  $F^{-1}$ . Let  $Y = F^{-1}(U)$ , then Y has the same distribution as X.

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Continuous rv - Inverse Transform Method Let X be a continuous random variable with cdf F(x) which possesses an inverse  $F^{-1}$ . Let  $Y = F^{-1}(U)$ , then Y has the same distribution as X.

**Proof** It is sufficient to show that Y has the same cdf as X. Let F and  $F_Y$  denote the cdfs of X and Y respectively. Then

$$F_{Y}(x) = P(Y \le x) = P(F^{-1}(U) \le x)$$
  
=  $P(F(F^{-1}(U) \le F(x)))$   
=  $P(U \le F(x)) = F(x)$ 

where the third-to-last equality follows from the fact that F is strictly increasing.

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Sampling from standard distributions

Uniform:  $X \sim U(a, b)$ : X = a + U(b - a)

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Sampling from standard distributions

Uniform: 
$$X \sim U(a, b)$$
:  $X = a + U(b - a)$   
Exponential:  $X \sim Exp(\lambda)$ :  $X = -\ln(U)/\lambda$   
Weibull:  $X \sim W(\alpha, \beta, \nu)$ :  $X = \nu + \alpha [-\ln(U)]^{1/\beta}$   
Cauchy:  $X \sim C(\mu, \sigma^2)$ :  $X = \mu + \sigma \tan \pi (U - \frac{1}{2})$ 

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Empirical CDF Suppose  $x_1 \le x_2 \le ... \le x_n$  is a collection of  $n$  data  
points, with  $x_i \in [a, \infty]$  for some  $a \in \mathbb{R}$ . Then the empirical CDF  $F(x)$   
with linear interpolation is defined as follows:

1. Given  $x \in \{x_1, \ldots, x_n\}$ , let *i* be the largest index for which  $x = x_i$  then  $F(x) = \frac{i}{n}$ 

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- 2. F(x) = 0 for all  $x \le a$
- 3. F(x) = 1 for all  $x \ge x_n$

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Weibull:  $X \sim W(\alpha, \beta, \nu)$ :  $X = \nu + \alpha[-\ln(U)]^{1/\beta}$   
Cauchy:  $X \sim C(\mu, \sigma^2)$ :  $X = \mu + \sigma \tan \pi (U - \frac{1}{2})$   
Empirical CDF Suppose  $x_1 \le x_2 \le ... \le x_n$  is a collection of  $n$  data  
points, with  $x_i \in [a, \infty]$  for some  $a \in \mathbb{R}$ . Then the empirical CDF  $F(x)$   
with linear interpolation is defined as follows:

- 1. Given  $x \in \{x_1, \ldots, x_n\}$ , let *i* be the largest index for which  $x = x_i$  then  $F(x) = \frac{i}{n}$
- 2. F(x) = 0 for all  $x \le a$
- 3. F(x) = 1 for all  $x \ge x_n$

4. if 
$$x \in (a, x_1)$$
 then  $F(x) = \frac{F(x_1)}{x_1 - a}(x - a)$ 

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Sampling from standard distributions

Uniform: 
$$X \sim U(a, b)$$
:  $X = a + U(b - a)$   
Exponential:  $X \sim Exp(\lambda)$ :  $X = -\ln(U)/\lambda$   
Weibull:  $X \sim W(\alpha, \beta, \nu)$ :  $X = \nu + \alpha[-\ln(U)]^{1/\beta}$   
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- 2. F(x) = 0 for all  $x \le a$
- 3. F(x) = 1 for all  $x \ge x_n$
- 4. if  $x \in (a, x_1)$  then  $F(x) = \frac{F(x_1)}{x_1 a}(x a)$
- 5. if  $x \in (x_i, x_{i+1})$  then  $F(x) = F(x_i) + \frac{(x-x_i)[F(x_{i+1}) F(x_i)]}{x_{i+1} x_i}$

Sampling from empirical CDFs Procedure for sampling a value from an empirical CDF F(x)

- 1. Sample from U
- 2. if U = 0 return a
- 3. else if  $U = F(x_i)$  for some  $1 \le i \le n$  then return  $x_i$
- 4. else if  $U < F(x_1)$  then return

$$a+(x_1-a)rac{U}{F(x_1)}$$

5. else if  $F(x_i) < U < F(x_{i+1})$  then return

$$x_i + (x_{i+1} - x_i) \frac{U - F(x_i)}{F(x_{i+1}) - F(x_i)}$$

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