## Extreme Value Theory in Civil Engineering

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## **Preliminaries: Return period**

- IID random variables  $\{X_1, X_2, X_3, ...\}$  with CDF  $F_X$ 
  - Occurrence (or success) =  $\{X_i > x_p\}$  in *i*<sup>th</sup> trial
  - $p = P\{success\} = P\{X_i > x_p\} = 1 F_X(x_p)$
  - $-x_p$  = level corresponding to exceedance probability p
- Sequence of independent and identical Bernoulli trials:
  - Geometric random variable
  - Time between successive occurences is random
  - "Mean Return Period" associated with  $x_p$  is 1/p (in units of trial time interval)

## **Preliminaries: Characteristic value**

- IID random variables  $\{X_1, X_2, X_3, ...\}$  with CDF  $F_X$ 
  - Success =  $\{X_i > x_p\}$  in *i*<sup>th</sup> trial
  - $p = P\{success\} = P\{X_i > x_p\} = 1 F_X(x_p)$
  - $-x_p$  = level corresponding to exceedance probability p
- *n* repeated trials Binomial random variable
  - Mean number of occurrences = np
- $x_n$  = Characteristic value of  $X_i$ 
  - if mean number of occurrences,  $np(x_n) = 1$
  - that is,  $1 F_X(x_n) = 1/n$

## **Motivation: Time-dependent reliability**

First Passage  $T = \inf [t : C(t, \underline{x}) < D(t, \underline{x}), t \ge 0, \underline{x} \in \Omega]$ Time:



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# **Design issues**

- Maximum load
- Minimum capacity
- Associated uncertainties
- Data driven

• Time-dependent failure probability

$$P_f(t) = P[R(\tau) - D - L(\tau) \le 0 \quad \text{for any} \quad \tau \in [0, t]]$$

• Simplification

$$P_f(t) = P[R_e - D - L_{\max,t} \le 0]$$

• Need to estimate  $L_{\max,t}$ 

# **Maximum live load estimation**

• Maximum live load

 $L_{\max,t} = \max\{L_1, L_2, ..., L_{N(t)}\}$ 

• Distribution function

$$F_{L_{\max,t}}(l) = P[L_1 \le l, L_2 \le l, ..., L_{N(t)} \le l]$$

- Simplifications
  - independence
  - stationarity

$$F_{L\max,t}(l) = [F_L(l)]^{N(t)}$$

# Extreme value theory: problem statement

- Sequence of random variables  $\{X_i\}$
- What are the limiting forms of  $Z_n$  and  $W_n$ ?

$$Z_{n} = \max(X_{1}, X_{2}, ..., X_{n}) \sim H_{n}$$
$$W_{n} = \min(X_{1}, X_{2}, ..., X_{n}) \sim L_{n}$$

- n unknown
- Degeneracy of limit distributions: n infinite
- Nature of population distributions,  $F_i$
- Dependence in the sequence
- Non-stationarity of the sequence

# IID (classical) case

- $\{X_i\}$  is an IID sequence -  $X_i$  and  $X_j$  are independent for  $i \neq j$ 
  - $-F_i = F$  are same for all i
- $H_n(x)$  and  $L_n(x)$  are degenerate distributions

$$P[Z_n \le x] = H_n(x) = F^n(x)$$
$$P[W_n \le x] = L_n(x) = 1 - (1 - F(x))^n$$

$$\lim_{n \to \infty} H_n(x) = \begin{cases} 0, x < \omega(F) \\ 1, \text{ otherwise} \end{cases}$$

 $\omega(F) =$  upper end point of F

$$\lim_{n \to \infty} L_{(n)}(x) = \begin{cases} 0, x \le \alpha(F) \\ 1, \text{ otherwise} \end{cases}$$

a(F) =lower end point of F

## Code for CDF of Xmax vs. n

- %This program generates a sequence of n IID exponential RVs and stores it maximum, xmax
- %It then repeats the process mct times
- %The distribution of xmax is plotted
- %The plot is repeated for a different n
- clear all;
- mct=1000;
- n=input('give n\n');
- for mcti=1:mct,
- for i=1:n,

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x(i)=-log(rand);
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- end
  - xmax(mcti)=max(x);

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freq(mcti)=mcti/(mct+1);
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- end
- xmaxsorted=sort(xmax);
- loglog(xmaxsorted,freq);
- sizes=['n=' num2str(n)];
- hold on;
- text(xmaxsorted(mct/2),.5,[sizes],'BackgroundColor',[1 1 1],'EdgeColor','black');
- axis([0,100,0,1]);

## Example: degeneracy for large n



• Can we normalize  $Z_n$  and help matters?

$$\lim_{n \to \infty} P\left[\frac{Z_n - a_n}{b_n} \le x\right] = \lim_{n \to \infty} H_n(a_n + b_n x) = H(x)$$

- Issues
  - How many possible forms for *H*?
  - How does H depend on F?
  - How to find  $a_n$  and  $b_n$
  - What is the speed of convergence?

• Can we normalize  $W_{(n)}$  and help matters?

$$\lim_{n \to \infty} P\left[\frac{W_{(n)} - a_n}{b_n} \le z\right] = \lim_{n \to \infty} L_{(n)}(a_n + b_n z) = L(z)$$

- Issues
  - How many possible forms for L ?
  - How does L depend on F?
  - How to find  $a_n$  and  $b_n$
  - What is the speed of convergence?

### Normalization

As  $n \to \infty$ , Instead of  $F_{\max}(x) = F_X^n(x)$ Look at  $F_X^n(x_n)$  where  $x_n = f(x, n)$ Simplest form for  $x_n$ :  $x_n = a_n + b_n x$ 

### **Normalization example**

$$X_{i} \sim \text{Exponential(1)}$$

$$F_{\max}(x) = F_{X}^{n}(x) = (1 - \exp(-x))^{n}$$
As  $n \to \infty$ ,  
Replace  $x \leftarrow \alpha(x - u) + \ln(n)$   
Obtain:  $F_{\max}(x) = \lim_{n \to \infty} \left(1 - \frac{\exp(-\alpha(x - u))}{n}\right)^{n}$   
 $= \exp(-\exp(-\alpha(x - u)))$ 

### Code for CDF of normalized Xmax vs. n

- %This program generates n IID Exponentials and stores the maximum xmax.
- %xmax is then recentered and scaled as a function of n
- %The process is repeated n times
- clear all;
- mct=1000;
- n=input('give n\n');
- scale=input('give scale\n');
- shift=input('give shift\n');
- if n==1,scale=1;shift=0;end
  - for mcti=1:mct,
    - for i=1:n,
      - $x(i) = -\log(rand);$
  - end
  - xmax(mcti)=(max(x)-log(n)+shift)/scale;
    - freq(mcti)=mcti/(mct+1);
- end
- xmaxsorted=sort(xmax);
- loglog(xmaxsorted,freq);
- sizes=['n=' num2str(n)];
- hold on;
- text(xmaxsorted(mct/2),.5,[sizes],'BackgroundColor',[1 1 1],'EdgeColor','black');
- axis([0,100,0,1]);



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# Limit distributions in IID case

- There are only three types of non-degenerate distributions *H*(*x*) for maxima
- There are only three types of non-degenerate distributions *L*(*x*) for minima
- Necessary and sufficient conditions exist for *F*(*x*) to yield above max or min distributions
  - Note: Two distributions *F* and *G* are of the same "type", if F(x) = G(ax+b) where *a*, *b* are constants

#### **Generalized EV distribution for maxima**

In IID case, H(z) must be of the same type as:

$$H_c(z) = \exp\left[-(1+cz)^{-1/c}\right], \ 1+cz > 0$$

 $c = 0 \Rightarrow \text{Gumbel (Type I) distribution:} \quad H_G(z) = e^{-e^{-z}}, -\infty < z < \infty$   $c > 0 \Rightarrow \text{Frechet (Type II) distribution:} \quad H_F(z) = \begin{cases} e^{-z^{-\gamma}}, z > 0\\ 0, z \le 0 \end{cases}$   $c < 0 \Rightarrow \text{Weibull (Type III) distribution:} \quad H_W(z) = \begin{cases} 1, & z > 0\\ e^{-(-z)^{-\gamma}}, z \le 0 \end{cases}$ 

where  $\gamma = |1/c|$ 

### **Generalized EV distribution for minima**

In IID case, L(z) must be of the same type as:

$$L_{c}(z) = 1 - \exp\left[-(1 - cz)^{-1/c}\right], \ 1 - cz > 0$$

 $c = 0 \Rightarrow \text{Gumbel (Type I) distribution: } L_G(z) = 1 - e^{-e^z}, -\infty < z < \infty$   $c > 0 \Rightarrow \text{Frechet (Type II) distribution: } L_F(z) = \begin{cases} 1 - e^{-(-z)^{-\gamma}}, z \le 0\\ 1, z > 0 \end{cases}$   $c < 0 \Rightarrow \text{Weibull (Type III) distribution: } L_W(z) = \begin{cases} 0, z < 0\\ 1 - e^{-z^{\gamma}}, z \ge 0 \end{cases}$ 

where  $\gamma = |1/c|$ 

#### **Domains of attraction for maxima**

(A) 
$$\lim_{t \to \infty} \frac{1 - F(tx)}{1 - F(t)} = x^{-\gamma}, \quad \gamma > 0$$
  
(B) 
$$\int_{a}^{\omega(F)} (1 - F(x)) dx < \infty, \text{ any finite } a, \quad \alpha(F), \quad \omega(F) = \text{ lower and upper end points of } F$$
  
(C) 
$$\lim_{t \to \omega(F)} \frac{1 - F(t + xR(t))}{1 - F(t)} = e^{-x}, \qquad R(t) = E(X - t \mid X > t), \quad t > \alpha(F)$$

•  $F \in D(H_F)$  if and only if  $\omega(F) = \infty$  and (A) holds for F  $a_n = 0, b_n = \inf(x : 1 - F(x) \le 1/n)$ •  $F \in D(H_W)$  if and only if  $\omega(F) < \infty$  and (A) holds for  $F^*(x) = F(\omega(F) - 1/x)$   $a_n = \omega(F), b_n = \omega(F) - \inf(x : 1 - F(x) \le 1/n)$ •  $F \in D(H_G)$  if and only if  $\omega(F) = \infty$  and (B), (C) hold

$$a_n = \inf(x: 1 - F(x) \le 1/n), b_n = R(a_n)$$

#### **Domains of attraction for minima**

(A) 
$$\lim_{t \to \infty} \frac{F(tx)}{F(t)} = x^{-\gamma}, \qquad \gamma > 0$$
  
(B) 
$$\int_{\alpha(F)}^{a} F(x) dx < \infty, \text{ any finite } a, \quad \alpha(F), \omega(F) = \text{ lower and upper end points of } F$$
  
(C) 
$$\lim_{t \to \alpha(F)} \frac{F(t + xr(t))}{F(t)} = e^{x}, \qquad r(t) = E(t - X \mid X < t), t > \alpha(F)$$

- $F \in D(L_F)$  if and only if  $\alpha(F) = -\infty$  and (A) holds for F $c_n = 0, d_n = \sup(x : F(x) \le 1/n)$
- $F \in D(L_w)$  if and only if  $\alpha(F) > -\infty$  and (A) holds for  $F^*(x) = F(\alpha(F) - 1/x), x < 0$

 $c_n = \alpha(F), d_n = \sup(x : F(x) \le 1/n) - \alpha(F)$ 

•  $F \in D(L_G)$  if and only if (B), (C) hold

 $c_n = \sup(x : F(x) \le 1/n), d_n = r(c_n)$ 



- Cauchy
- Uniform
- Exponential
- Rayleigh

#### **Examples**

Limit distribution for minima from Cauchy parent

$$F(x) = \frac{1}{2} + \frac{\arctan(x)}{\pi}; \quad -\infty < x < \infty$$
  
Check Eqn (A) for minima:  
$$\lim_{t \to \infty} \frac{F(tx)}{f(t)} = \lim_{t \to \infty} \frac{\frac{1}{2} + \frac{\arctan(tx)}{\pi}}{\frac{1}{2} + \frac{\arctan(tx)}{\pi}}$$
$$= \lim_{t \to \infty} \frac{\frac{x}{1 + (tx)^2}}{\frac{1}{1 + t^2}}$$
$$= \lim_{t \to \infty} \frac{x(1 + t^2)}{1 + t^2 x^2} = x^{-1}$$

That is,  $\gamma = -1$ , and Cauchy lies in the domain of attraction of Frechet for minima. The normalizing constants are:

$$c_n = 0$$
$$d_n = \tan\left[\pi\left(\frac{1}{2} - \frac{1}{n}\right)\right]$$

#### **Examples**

• Limit distribution for maxima from uniform parent

The complemetary CDF of the uniform is:

$$F^{*}(x) = 1 - \frac{1}{x}, x \le 1$$
$$\lim_{t \to \infty} \frac{1 - F^{*}(tx)}{1 - F^{*}(t)} = \lim_{t \to \infty} \frac{\frac{1}{tx}}{\frac{1}{t}} = x^{-1}$$

Since  $\gamma = 1$ , uniform gives rise to Weibull maxima.

The normalizing constants are:

$$a_n = \omega(F) = 1$$
  
$$b_n = \omega(F) - F^{-1} \left( 1 - \frac{1}{n} \right)$$
  
$$= 1 - 1 + \frac{1}{n} = \frac{1}{n}$$



• Limit distribution for minima from exponential parent

$$F^*(x) = 1 - \exp\left(\frac{1}{x}\right)$$
$$\lim_{t \to -\infty} \frac{F^*(tx)}{F^*(t)} = \lim_{t \to -\infty} \frac{1 - \exp\left(\frac{1}{tx}\right)}{1 - \exp\left(\frac{1}{t}\right)} = \lim_{t \to -\infty} \frac{\exp\left(\frac{1}{tx}\right)\left(\frac{1}{t^2}\right)\left(\frac{1}{x}\right)}{\exp\left(\frac{1}{t}\right)\left(\frac{1}{t^2}\right)} = x^{-1}$$

Since  $\gamma = 1$ , exponential gives rise to Weibull minima.  $c_n = \alpha(F) = 0$  $d_n = F^{-1}\left(\frac{1}{n}\right) - \alpha(F) = -\log\left(1 - \frac{1}{n}\right) \approx \frac{1}{n}$ 

## **Domains of attraction**

Domain of Attraction Type		
Distribution	For maximum	For minimum
Normal	Gumbel	Gumbel
Exponential	Gumbel	Weibull
Log-normal	Gumbel	Gumbel
Gamma	Gumbel	Weibull
Gumbel <sub>M</sub>	Gumbel	Gumbel
Gumbel <sub>m</sub>	Gumbel	Gumbel
Rayleigh	Gumbel	Weibull
Uniform	Weibull	Weibull
Weibull <sub>M</sub>	Weibull	Gumbel
Weibull <sub>m</sub>	Gumbel	Weibull
Cauchy	Frechet	Frechet
Pareto	Frechet	Weibull
Frechet M	Frechet	Gumbel
Frechet m	Gumbel	Frechet

M=for maximum m=for minimum

# Estimation of EV distribution parameters

- Block maxima probability plot
- Return period plot
- Problem: not all data are utilized

#### **Generalized Pareto Distribution**

**A** 1

Exceedances of X over high threshold uDefine: Y = X - u

$$G(y) = P[Y \le y | Y > 0] = 1 - [1 + (cy/a)]^{-1/c}$$
  
$$a > 0, 1 + (cy/a) > 0$$

same c as in GEV distribution